Last week: • Examples

- Definitions + 3.4.1
- More examples and 3.4.2 Spatial statistical dependency

This week: Simulations + more models (3.4.3 + revisit ++). Estimation

Extra text book: Statistical Analysis and Modelling of Spatial Point Patterns by Janine Illian, Antti Penttinen, Helga Stoyan, Dietrich Stoyan (2008) TODAY:

- Clustering Neymann-Scott processes
 - Cox processes
 - Log-Gaussian Cox process

Regularization • Gibbs processes

From Cressie and Wikle: 4.3 SPATIAL POINT PROCESSES

A spatial point process is a stochastic process governing the location of events (equivalently, points) $\{\mathbf{s}_i\}$ in some set $D_s \subset \mathbb{R}^d$, where the number of such events in D_s is also random (e.g., Diggle, 2003). In the simplest case, the

and we only consider:

Only *simple* spatial point processes in \mathbb{R}^d (i.e., almost surely, either no event or a single event occurs at any point) will be considered. We characterize the

Data

Number of points: $Z(D_s) = m$ Locations: $\{s_1, s_2, \dots, s_m\}$

from Illian et al 2008:

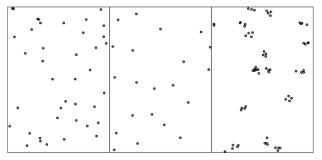


Figure 1.1 Three simulated point patterns: (left) random, (centre) regular, (right) clustered.

Ex. Random: Homogenous Poisson point process (HPPP)

An important, but simple, example is the Poisson point process Z, for which

$$Z(A)|\lambda^{o} \sim Poi(\lambda^{o}|A|), \qquad A \subset D_{s}, \qquad (4.159)$$

where $\lambda^o > 0$ is a parameter of the Poisson point process, and recall that |A| is the *d*-dimensional volume of *A*. More details on this spatial point process are given in Section 4.3.1.

Sampling from Poisson point process

Trees in $1km \times 1km$ domain, with intensity $\lambda^0 = 15 trees/km^2$

- Sample $m \sim Pos(15)$
- for *i* = 1 : *m*
 - sample location randomly in domain $\{s_i\}$
- end

Clustering, how to sample/model?

• ...

February 14, 2017, Ingelin Steinsland TMA 4250:Point Processes

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Data model: $[Z|Y, \theta]$ Process model: $[Y|\theta]$ Parameter model: $[\theta]$

Cox process

Data model: $[Z|\lambda] \sim Pos(\int_A \lambda(d)dx)$ Process model: $\lambda(s)$ (Cox process)

from Note

| Name of Process | Clusters | Parents |
|-----------------------------|---------------------------|------------------|
| Independent cluster process | general | general |
| Poisson cluster process | general | Poisson |
| Cox cluster process | Poisson | general |
| Neyman-Scott process | Poisson | Poisson |
| Matérn cluster process | Poisson (uniform in ball) | Poisson (homog.) |
| Modified Thomas Process | Poisson (Gaussian) | Poisson (homog.) |

Table 2: Nomenclature for independent cluster processes.

Sampling Neymann-Scott

from Note

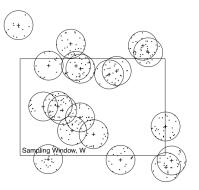


Figure 1: Edge effects for a Matérn process. Parent process is homogeneous Poisson with a constant rate of 20 in a 20×20 square field. The cluster processes are homogeneous Poisson centered at the parents (+) with a radius of 3 and a constant rate of 10.

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ds denote a small region located at s with volume |ds|. Then the *first-order intensity* function of the point process $Z(\cdot)$ is defined as

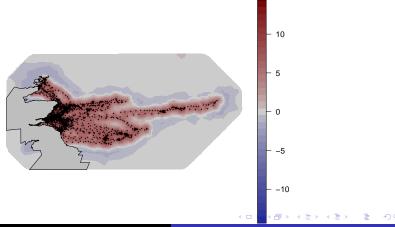
$$\lambda(\mathbf{s}) \equiv \lim_{|d\mathbf{s}| \to 0} E(Z(d\mathbf{s}))/|d\mathbf{s}|, \qquad \mathbf{s} \in D_s \,, \tag{4.160}$$

provided the limit exists. Hence,

$$E(Z(A)) = \int_{A} \lambda(\mathbf{s}) \, d\mathbf{s}, \qquad A \subset D_{s}. \tag{4.161}$$

Estimation of intensity

We often want to estimate the $\lambda(s)$ from data (points). Seals (from log-Gaussian Cox model):



Kernel estimation of intensity

SPATIAL POINT PROCESSES

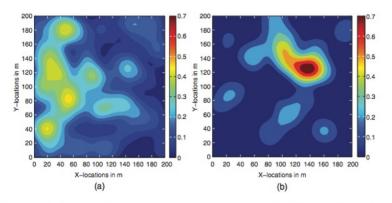


Figure 4.16 (a) Contour plot of estimated first-order intensity of $m_A = 271$ adult longleaf pine trees in the 4-ha study area in the Wade Tract; bandwidth b = 30 m. (b) Contour plot of estimated first-order intensity of $m_S = 159$ subadult longleaf pine trees in the same study area as (a); bandwidth b = 30 m.

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207

2nd order intensity and Pair-correlation function

2nd order intensity

$$\lambda_2(s,x) = \lim_{|ds| \to 0, |dx| \to 0} \frac{E(Z(ds)Z(dx))}{|ds||dx|}$$

Pair correlation function

$$g(s,x) = \frac{\lambda_2(x,s)}{\lambda(s)\lambda(x)}$$

For HPPP: $\lambda_2(s,x) = \lambda(s)\lambda(x)$ and g(s,x) = 1

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Data model: $[Z|Y, \theta]$ Process model: $[Y|\theta]$ Parameter model: $[\theta]$

Log Gaussian Cox process (LGCP)

Data model: Conditional on $\lambda(\cdot)$, the point process $Z(\cdot)$ is an inhomogeneous Poisson point process with intensity $\lambda(\cdot)$. Process model: Conditional on β and $C_Y(\cdot, \cdot)$, $Y(\cdot) \equiv \log \lambda(\cdot)$ is a Gaussian

process with the following properties:

 $E(Y(\mathbf{s})) = \mathbf{x}(\mathbf{s})'\boldsymbol{\beta}$ and $C_Y(\mathbf{s}, \mathbf{x}) \equiv \operatorname{cov}(Y(\mathbf{s}), Y(\mathbf{x}))$.

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 $\lambda_2(s, u)$ for Log Gaussian Cox Process (LGCP)

From page 214 Cressie & Wikle.

$$1$$

$$E(Z(ds)Z(dx)) = E\{E(Z(ds)Z(dx)|\lambda(\cdot))\}$$

$$2 = E\{\lambda(ds)\lambda(dx)\}$$

$$3 = E(e^{Y(s)} \cdot e^{Y(x)})|ds||dx|, \quad (4.182)$$

 $\lambda_2(\mathbf{s}, \mathbf{x}) = \exp\{\mu_Y(\mathbf{s}) + \mu_Y(\mathbf{x}) + C_Y(\mathbf{s}, \mathbf{s})/2 + C_Y(\mathbf{x}, \mathbf{x})/2 + C_Y(\mathbf{s}, \mathbf{x})\}.$ (4.183)

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Intensity functions

1 order intensity:

$$\lambda_1(s) = \lim_{|ds| \to 0} \frac{E(Z(ds))}{|ds|}$$

2. order intensity:

$$\lambda_2(s_1, s_2) = \lim_{|ds_1| \to 0, |ds_2| \to 0} \frac{E(Z(ds_1)Z(ds_2))}{|ds_1||ds_2|}$$

*m*th order intersity

$$\lambda_m(s_1, s_2) = \lim_{\dots \to 0} \frac{E(Z(ds_1)Z(ds_2)\dots Z(ds_m))}{|ds_1||ds_2|\dots |ds_m|}$$

from Note. $\lambda^0 = \beta = 100$

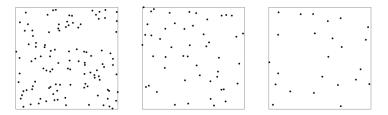


Table 3: Realizations of Strauss processes on $S = [0, 1]^2$. Here $\beta = 100$, R = 0.1 and $\gamma = 1.0, 0.5, 0.0$ from left to right. Processes generated from the function rStrauss in the R package spatstat by Adrian Baddeley.